

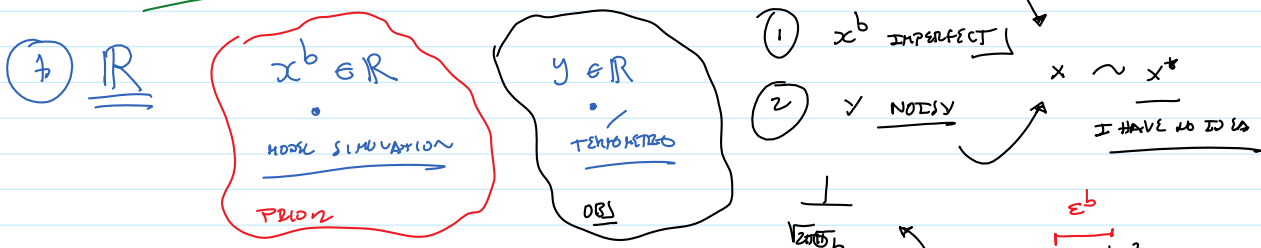
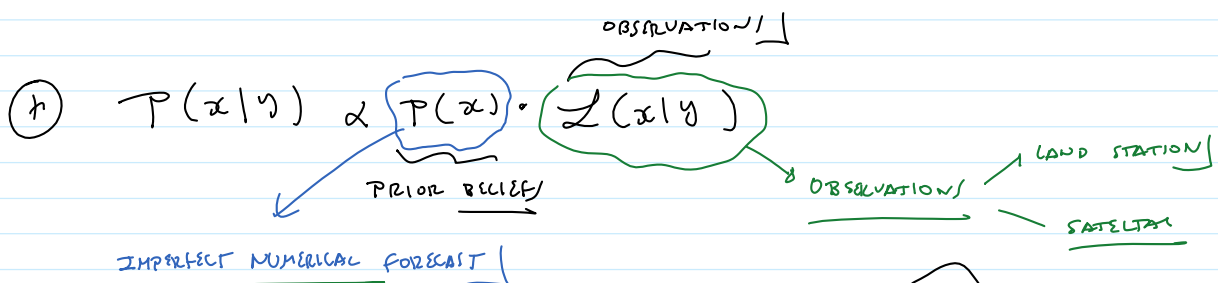
Data Assimilation - Bayes' Theorem

miércoles, 12 de agosto de 2020 01:44 p. m.

$$P(x|y) \propto \underbrace{P(x)}_{\text{PRIOR}} \cdot \underbrace{\mathcal{L}(x|y)}_{\text{LIKELIHOOD}} \quad \left. \begin{matrix} P(x|y) \\ P(x) \\ \mathcal{L}(x|y) \end{matrix} \right\} \text{PDF}$$

$$P(x|y) = \frac{P(x) \cdot \mathcal{L}(x|y)}{P(y)}$$

$$= \frac{1}{P(y)} \cdot P(x) \cdot \mathcal{L}(x|y) \propto \underbrace{P(x)}_{\text{PRIOR BELIEF}} \cdot \underbrace{\mathcal{L}(x|y)}_{\text{PRIOR OBSERVATION}}$$



ASSUMPTIONS:

1) $x \sim \mathcal{N}(x^b, \sigma_b^2) \quad P(x) \propto e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2}}$

2) $y \sim \mathcal{N}(x, \sigma_o^2) \quad \mathcal{L}(x|y) \propto e^{-\frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$

$$\varepsilon^b = x - x^b \sim \mathcal{N}(0, \sigma_b^2) \quad \varepsilon^o = y - x \sim \mathcal{N}(0, \sigma_o^2)$$

3) $\text{cov}(\varepsilon^b, \varepsilon^o) = 0$ NO CORRELATED ERRORS

$$P(x|y) \propto P(x) \cdot \mathcal{L}(x|y) = e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2}} \cdot e^{-\frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$$

MEAN, VAR, MEAN, VAR

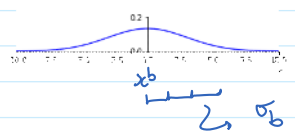
$$P(x|y) = ?$$

$$P(x|y) \propto e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2}} \cdot e^{-\frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$$

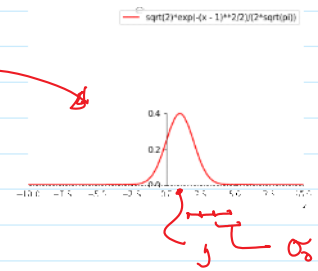
$$P(x|y) \propto P(x) \cdot \mathcal{L}(x|y) \propto e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2}} \cdot e^{-\frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$$

$$= e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2} - \frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$$

$$= e^{-\mathcal{J}(x)} \quad \mathcal{J}(x) = \frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2} + \frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}$$

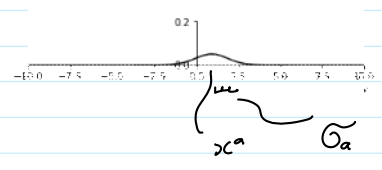


$N(x^b, \sigma_b^2)$



$$\exp(-x^2/2) \exp(-(x-1)^2/2)$$

$N(x, \sigma_o^2)$



$N(x^a, \sigma_a^2)$

MAXIMUM
A
MAP - POSTERIOR
ESTIMATE

$$x^a = \arg \max_x P(x|y)$$

(+)

$$P(x|y) \propto e^{-\mathcal{J}(x)}$$

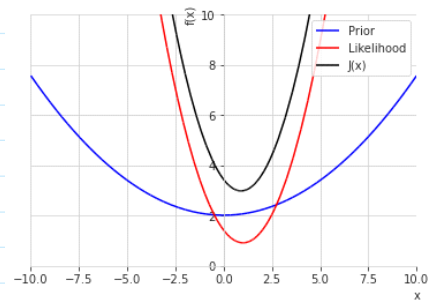
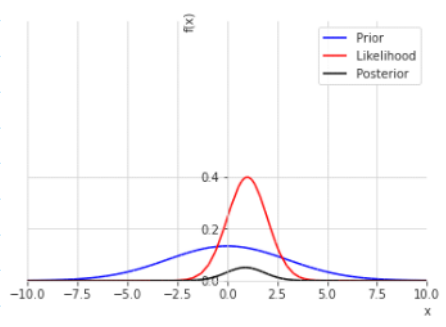
$x^a = ?$ } ANALYSIS

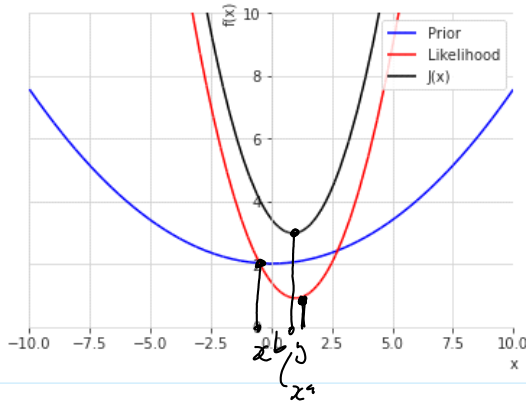
$\sigma_a = ?$

$$P(x|y) \propto e^{-\frac{1}{2} \mathcal{J}(x)} \quad \mathcal{J}(x) = \frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2} + \frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}$$

=

$$x^a = \arg \max_x P(x|y) \quad x^a = \arg \min_x \mathcal{J}(x)$$





$$\textcircled{\phi} \quad J(x) = \frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2} + \frac{1}{2} \frac{(y-x)^2}{\sigma_o^2} \quad J'(x) = \frac{x-x^b}{\sigma_b^2} - \frac{y-x}{\sigma_o^2}$$

$$J'(x) = \frac{x}{\sigma_b^2} - \frac{x^b}{\sigma_b^2} - \frac{y}{\sigma_o^2} + \frac{x}{\sigma_o^2} = \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x - \left[\frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2} \right]$$

$$J'(x^a) = 0 \Rightarrow x^a = \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right)^{-1} \left(\frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2} \right)$$

$$\left[\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right]^{-1} = \frac{\sigma_o^2 \cdot \sigma_b^2}{\sigma_b^2 + \sigma_o^2} \quad x^a = \left[\frac{\sigma_o^2 \cdot \sigma_b^2}{\sigma_b^2 + \sigma_o^2} \right] \left[\frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2} \right]$$

$$\Rightarrow x^a = \frac{\sigma_o^2}{\sigma_b^2 + \sigma_o^2} \cdot x^b + \frac{\sigma_b^2}{\sigma_b^2 + \sigma_o^2} \cdot y$$

$\alpha_1 + \alpha_2 = 1$
 $\alpha_1 \in [0, 1]$
 $\alpha_2 \in [0, 1]$

CONVEX COMBINATION

$$x^a = \alpha x^b + (1-\alpha) y$$

WHAT MODEL SAYS

WHAT OBSERVATION SAYS

$$\alpha = \frac{\sigma_o^2}{\sigma_b^2 + \sigma_o^2}$$

$$\textcircled{+} \quad P(x|y) \propto \underbrace{e^{-\frac{1}{2} \frac{(x-x^a)^2}{\sigma_a^2}}}_{\text{SAY?}} \quad \sigma_a = ?$$

$$P(x|y) \propto P(x) \cdot \mathcal{L}(x|y) \propto e^{-\frac{1}{2} \frac{(x-x^b)^2}{\sigma_b^2} - \frac{1}{2} \frac{(y-x)^2}{\sigma_o^2}}$$

$$\propto e^{-\frac{1}{2} \left[\frac{(x-x^b)^2}{\sigma_b^2} + \frac{(y-x)^2}{\sigma_o^2} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\frac{x^2}{\sigma_b^2} - \frac{2xx^b}{\sigma_b^2} + \frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} - \frac{2yx}{\sigma_o^2} + \frac{x^2}{\sigma_o^2} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\frac{x^2}{\sigma_b^2} - \frac{2xx^b}{\sigma_b^2} - \frac{2yx}{\sigma_o^2} + \frac{x^2}{\sigma_o^2} + \frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^2 - \left(\frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2} \right) 2x + \frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} \right]}$$

$$\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^a = \frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2}$$

$$\propto e^{-\frac{1}{2} \left[\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^2 - \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) 2xx^a + \frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^2 - \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) 2xx^a \right]} \cdot e^{-\frac{1}{2} \left[\frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^2 - \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) 2xx^a \right]} \underbrace{\left[\frac{x^{b2}}{\sigma_b^2} + \frac{y^2}{\sigma_o^2} \right]}_{\text{CONSTANT}}$$

PROP TO

$$\propto e^{-\frac{1}{2} \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) [x^2 - 2xx^a]} \cdot e^{-\frac{1}{2} \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) x^{a2}}$$

$$\propto e^{-\frac{1}{2} \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) (x^2 - 2xx^a + x^{a2})}$$

$$\propto e^{-\frac{1}{2} \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2} \right) (x-x^a)^2} \propto e^{-\frac{1}{2} \frac{(x-x^a)^2}{\left(\frac{\sigma_o^2 \cdot \sigma_b^2}{\sigma_o^2 + \sigma_b^2} \right)}}$$

1ST MOMENT: $x^a = \left(\frac{\sigma_o^2 \cdot \sigma_b^2}{\sigma_o^2 + \sigma_b^2} \right) \left(\frac{x^b}{\sigma_b^2} + \frac{y}{\sigma_o^2} \right)$

2nd MOMENT $\sigma_a^2 = \frac{\sigma_0^2 \cdot \sigma_b^2}{\sigma_0^2 + \sigma_b^2}$

⊕

EXAMPLE

LINEAR MODEL

$\phi_i(x) = x_i$

⊕

$y = \sum_{i=0}^N \alpha_i \cdot \phi_i(x) + \varepsilon$ $\varepsilon \sim \mathcal{N}(0, \sigma^2)$

⊕

$y_1 = \alpha_1 x_1^{(1)} + \alpha_2 x_2^{(1)} + \dots + \alpha_N x_N^{(1)} + \varepsilon_1$
 $y_2 = \alpha_1 x_1^{(2)} + \alpha_2 x_2^{(2)} + \dots + \alpha_N x_N^{(2)} + \varepsilon_2$
 \vdots
 $y_m = \alpha_1 x_1^{(m)} + \alpha_2 x_2^{(m)} + \dots + \alpha_N x_N^{(m)} + \varepsilon_m$

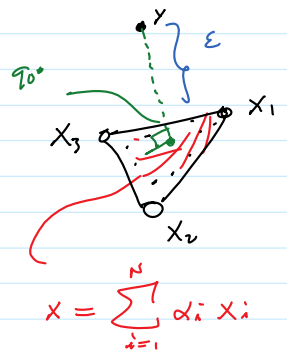
$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{bmatrix} = \begin{bmatrix} x_1^{(1)} & x_2^{(1)} & \dots & x_N^{(1)} \\ x_1^{(2)} & x_2^{(2)} & \dots & x_N^{(2)} \\ \vdots & \vdots & \ddots & \vdots \\ x_1^{(m)} & x_2^{(m)} & \dots & x_N^{(m)} \end{bmatrix} \cdot \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_m \end{bmatrix} + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_m \end{bmatrix}$$

$\text{cov}(\varepsilon_i, \varepsilon_j) = 0$
 $\text{cov}(\varepsilon_i, \varepsilon_j) = 0 \quad i \neq j$
 $\text{cov}(\varepsilon_i, \varepsilon_i) = \sigma^2 \quad i=j$

$y \quad X \quad \alpha \quad \varepsilon \in \mathbb{R}^{m \times 1}$

$\Rightarrow y = X\alpha + \varepsilon \quad \varepsilon \sim \mathcal{N}(0, \sigma^2 \cdot I)$

$$y = \sum_{i=1}^N \alpha_i x_i + \varepsilon$$
 $x_i \in \mathbb{R}^{m \times 1}$



$$\varepsilon(\alpha) = y - X\alpha$$

$y - X\alpha \sim \mathcal{N}(0, \sigma^2 \cdot I)$

$y - X\alpha \sim \mathcal{N}(0, \sigma^2 I) \equiv \mathcal{L}(\alpha | X, y) \propto e^{-\frac{1}{2} \|y - X\alpha\|_{(\sigma^2 I)}^2}$

$$P(x) \propto e^{-\frac{1}{2}(x-\mu)^T Q^{-1}(x-\mu)} = e^{-\frac{1}{2}\|x-\mu\|_{Q^{-1}}^2}$$

↳ KENSE OJT NORMA $\mathbb{R}^{n \times n}$

$$y - x\alpha \sim \mathcal{N}(0, \sigma^2 I) \equiv \mathcal{L}(\alpha | x, y) \propto e^{-\frac{1}{2}\|y - x\alpha\|_{(\sigma^2 I)^{-1}}^2}$$

⊕

$$P(\alpha | x, y) \propto P(\alpha) \cdot \mathcal{L}(x | y) \quad \sqrt{\mathbb{E}((y - x\alpha)(y - x\alpha)^T) = \sigma^2 I}$$

$$P(\alpha | x, y) \propto \mathcal{L}(x | y)$$

$$\propto e^{-\frac{1}{2}\|y - x\alpha\|_{(\sigma^2 I)^{-1}}^2} = e^{-J(\alpha)}$$

$$J(\alpha) = \frac{1}{2}\|y - x\alpha\|_{(\sigma^2 I)^{-1}}^2 = \frac{1}{2}(y - x\alpha)^T (\sigma^2 I)^{-1} (y - x\alpha)$$

$$= \frac{1}{2\sigma^2} (y - x\alpha)^T (y - x\alpha)$$

$$\alpha^* = \arg \max_{\alpha} P(\alpha | x, y) \equiv \alpha^* = \arg \min_{\alpha} J(\alpha)$$

$$\nabla_{\alpha} (J(\alpha)) = \nabla_{\alpha} \left(\frac{1}{2\sigma^2} (y - x\alpha)^T (y - x\alpha) \right)$$

$y \in \mathbb{R}^{m \times 1}$ $x \in \mathbb{R}^{m \times n}$ $\alpha \in \mathbb{R}^{n \times 1}$
 $z \in \mathbb{R}^{m \times 1}$ $z \in \mathbb{R}^{m \times 1}$

$z^T y = y^T z$

$$\nabla_{\alpha} (J(\alpha)) = \nabla_{\alpha} \left(\frac{1}{2\sigma^2} (y^T y - (x\alpha)^T y - y^T (x\alpha) + (x\alpha)^T (x\alpha)) \right)$$

$$= \nabla_{\alpha} \left(\frac{1}{2\sigma^2} (y^T y - 2(x\alpha)^T y + \alpha^T x^T x \alpha) \right)$$

$$= \nabla_{\alpha} \left(\frac{1}{2\sigma^2} (\alpha^T x^T x \alpha - 2\alpha^T x^T y + y^T y) \right)$$

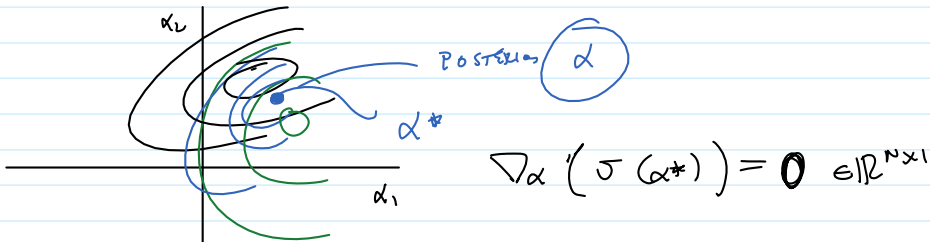
$$= \frac{1}{2\sigma^2} \left[\nabla_{\alpha} (\alpha^T x^T x \alpha) - 2 \nabla_{\alpha} (\alpha^T x^T y) + \nabla_{\alpha} (y^T y) \right]$$

⊕ $\mathbb{R}^{n \times 1}$

$$= \frac{1}{2\sigma^2} [2x^T x \alpha - 2x^T y] = [x^T x \alpha - x^T y] \cdot \frac{1}{\sigma^2}$$

$$= \frac{1}{2\sigma^2} [2X^T X \alpha - 2X^T y] = [X^T X \alpha - X^T y] \cdot \frac{1}{\sigma^2}$$

$$\nabla_{\alpha} (J(\alpha)) = \frac{1}{\sigma^2} \cdot [X^T X \alpha - X^T y] \quad \left| \begin{array}{l} \text{FIRST DERIVATIVE OF} \\ \nabla_{\alpha} (J(\alpha)) \end{array} \right.$$



$$\nabla_{\alpha} (J(\alpha^*)) = \mathbf{0} \in \mathbb{R}^{N \times 1}$$

$$\Rightarrow \nabla_{\alpha} (J(\alpha^*)) = \frac{1}{\sigma^2} [X^T X \alpha^* - X^T y] = \mathbf{0} \in \mathbb{R}^{N \times 1}$$

$$= \mathbf{0}$$

$$\alpha^* = (X^T X)^{-1} X^T y$$

∇ $P(\alpha | X, y) \propto e^{-\frac{1}{2} \|\alpha - \alpha^*\|^2} \binom{Z}{Z}^{-1}$

$P(\alpha | X, y) \propto e^{-\frac{1}{2\sigma^2} (y^T y - 2\alpha^T X^T y + (X\alpha)^T (X\alpha))}$

α^* $\propto e^{-\frac{1}{2\sigma^2} (\alpha^T X^T X \alpha - 2\alpha^T X^T y)}$

$$\propto e^{-\frac{1}{2\sigma^2} (\alpha^T X^T X \alpha - 2\alpha^T (X^T X) \alpha^*)}$$

$$\propto e^{-\frac{1}{2\sigma^2} (\alpha^T X^T X \alpha - 2\alpha^T X^T X \alpha^*)} \cdot e^{-\frac{1}{2\sigma^2} \alpha^{*T} X^T X \alpha^*}$$

$$\propto e^{-\frac{1}{2\sigma^2} (\alpha^T X^T X \alpha - 2\alpha^T X^T X \alpha^* + \alpha^{*T} X^T X \alpha^*)}$$

$$\propto e^{-\frac{1}{2\sigma^2} (\alpha^T X^T X \alpha - \alpha^T X^T X \alpha^* - \alpha^T X^T X \alpha^* + \alpha^{*T} X^T X \alpha^*)}$$

$$\alpha \cdot e^{-\frac{1}{2\sigma^2} \left(\alpha^T X^T X \alpha - \alpha^T X^T X \alpha^* - \alpha^T X^T X \alpha^* + \alpha^{*T} X^T X \alpha^* \right)}$$

$$\alpha \cdot e^{-\frac{1}{2\sigma^2} \left(\alpha^T X^T X [\alpha - \alpha^*] + \alpha^{*T} X^T X \alpha^* - \alpha^{*T} X^T X \alpha \right)}$$

$$\Rightarrow \alpha \cdot e^{-\frac{1}{2\sigma^2} \left[\alpha^T X^T X [\alpha - \alpha^*] - \alpha^{*T} X^T X [\alpha - \alpha^*] \right]}$$

$$\alpha \cdot e^{-\frac{1}{2\sigma^2} \left[(\alpha - \alpha^*)^T X^T X (\alpha - \alpha^*) \right]}$$

$$\alpha \cdot e^{-\frac{1}{2} \left[(\alpha - \alpha^*)^T \left[\frac{1}{\sigma^2} X^T X \right] (\alpha - \alpha^*) \right]}$$

$$\alpha \cdot e^{-\frac{1}{2} \left[(\alpha - \alpha^*)^T \left[\sigma^2 \cdot (X^T X)^{-1} \right] (\alpha - \alpha^*) \right]}$$

$$\alpha \cdot e^{-\frac{1}{2} \|\alpha - \alpha^*\|^2 \left(\frac{1}{\sigma^2} X^T X \right)^{-1}}$$

$$P(\alpha | X, y) \propto e^{-\frac{1}{2} \|\alpha - \alpha^*\|^2 \left(\frac{1}{\sigma^2} X^T X \right)^{-1}}$$

POSTERIOR DISTRIBUTION
POSTERIOR COVARIANCE

BAYESIAN INFERENCE

MAP = MEAN = MODIAN

$$P(x|y) \propto P(x) \cdot \mathcal{L}(x|y)$$

⇒ COMPUTE POSTERIOR PARAMETERS

↳ KERNEL → KERNEL POSTERIOR

- 1) COMPUTE MAP
- 2) COMPUTE COVARIANCE

$$P(x|y) \propto P(x) \cdot \mathcal{L}(x|y) \quad \mapsto \quad P(\alpha|x,y) \propto \mathcal{L}(\alpha|x,y)$$

